

Basics of Data Analysis

Measurement uncertainties / Error analysis

Slides developed and adapted by

Kirill Bolotin, Martin Weinelt, and Niclas Müller

- P. R. Bevington, and D. K. Robinson, *“Data Reduction and Error Analysis for the Physical Sciences”*
- P. Möhrke, and B.-U. Runge, *“Arbeiten mit Messdaten”*
- *“Guide to uncertainty propagation and Error analysis”* Stony Brook U.
- *“A beginner guide to uncertainty of measurement”* by Stephanie Bell
- Slides and python code examples will be uploaded to website

Basics of Data Analysis

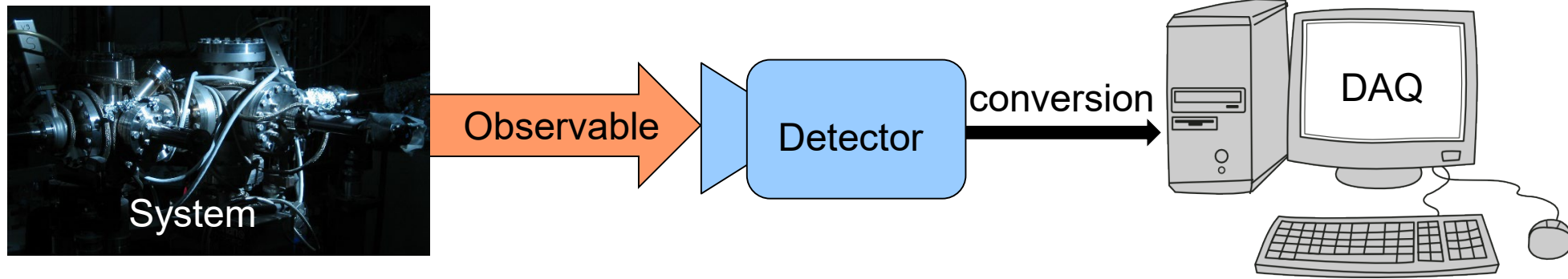
Measurement uncertainties / Error analysis

based on

- Slides by Kirill Bolotin and Martin Weinelt
- P. R. Bevington, and D. K. Robinson, “*Data Reduction and Error Analysis for the Physical Sciences*”
- P. Möhrke, and B.-U. Runge, “*Arbeiten mit Messdaten*”
- Check out also: T. Kampfrath, Grundlagen der Mess- und Labortechnik, Fri 2-4pm, Large Lect Hall

Always report measurement results with uncertainties!

Errors and uncertainties in experimental data



- Environmental
 - Wind, magnetic field, temperature
- Instrumental
 - Instrument accuracy limit
 - Fluctuating signal
- Observational
 - Misused instrument
- Theoretical
 - Error in equation
 - Effect not accounted for

- Types of errors and uncertainties
- Characterization of distributions
- Probability distributions

- Uncertainty analysis
- How to report and compare results
- Data fitting

Sources of uncertainty

Systematic error

- Reproducible inaccuracy
 - Faulty instrumentation
 - Neglected effect
-
- Getting rid of requires physics understanding!
 - Improving measurement equipment
 - Instrument calibration!
 - **Statistical analysis DOES NOT help!**

Random fluctuations

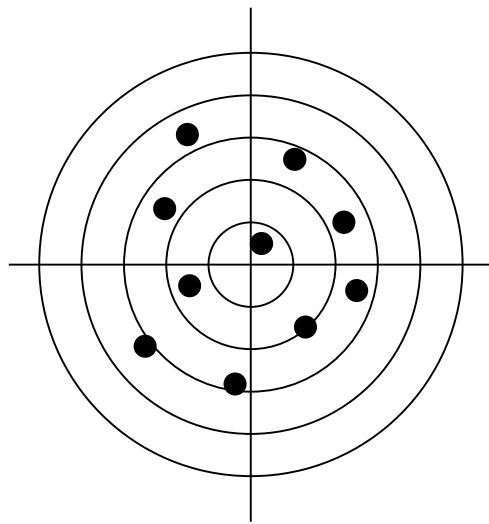
- Statistical variations
 - Limited precision
 - Fluctuating results
-
- Repeating experiments required!
 - **Statistical analysis helps!**

➤ **Goal: Estimate uncertainty to extract maximum of information!**

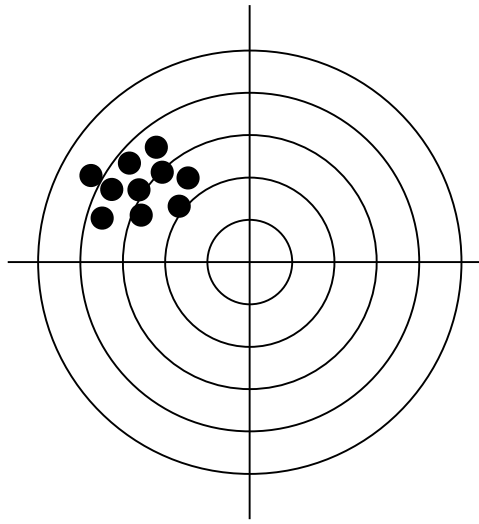
Sources of uncertainty

Accuracy: a measure how close a result is to the true value

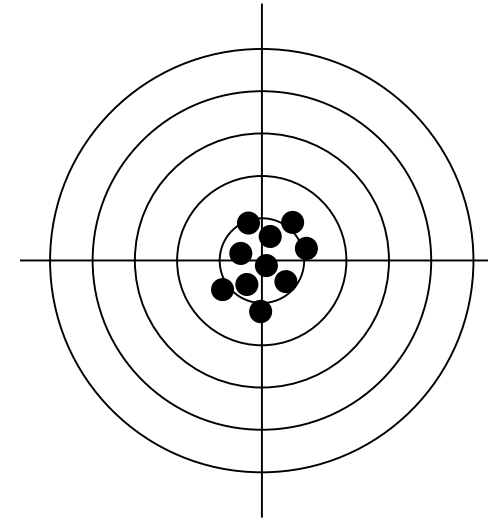
Precision: a measure how well a result has been determined without reference to the true value



accurate
but **imprecise**



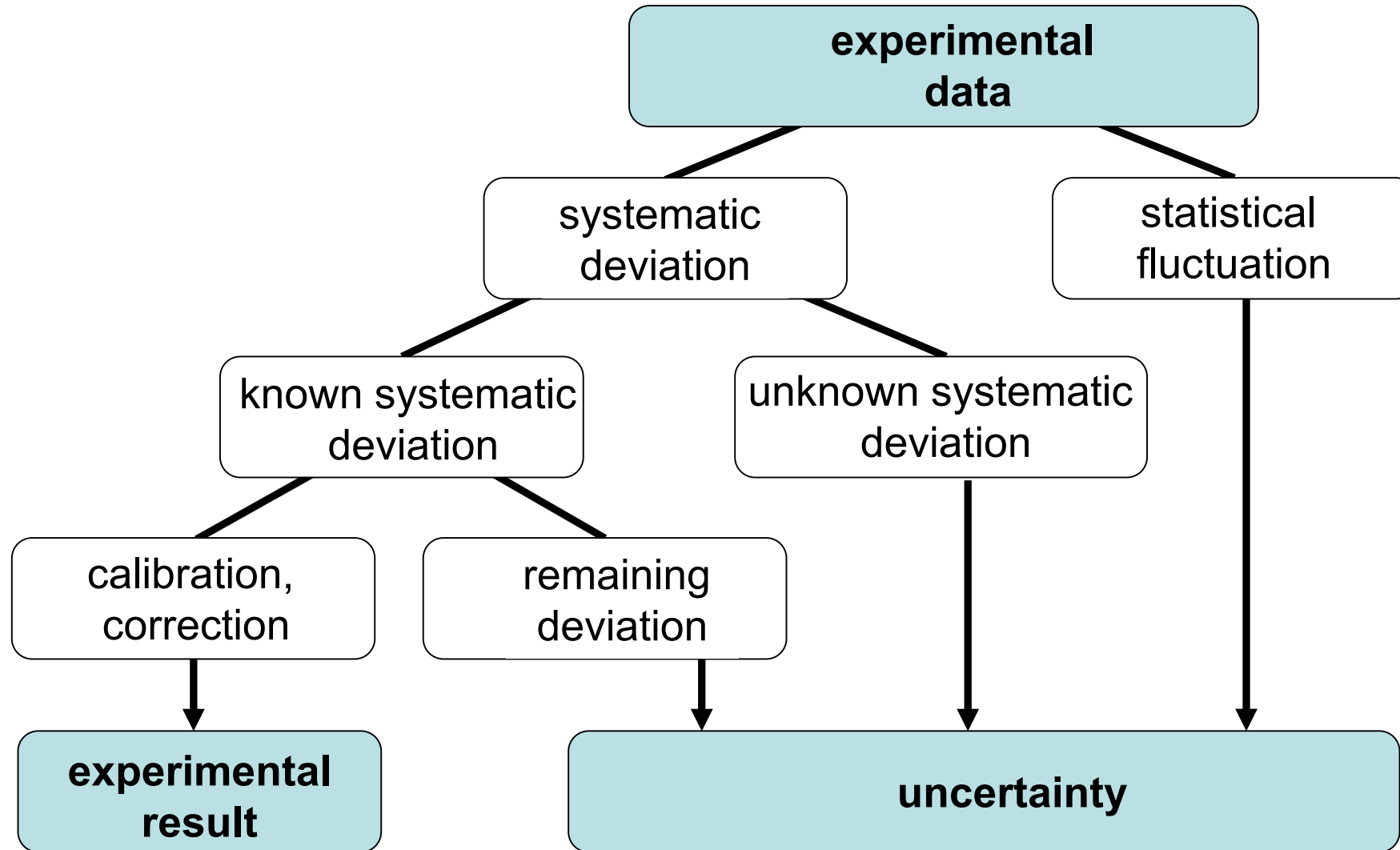
precise
but **inaccurate**



accurate and precise

Accuracy is mostly related to systematic deviations, *precision* mostly to statistics.

Work flow



modified from M. Hemla, QZ. 41, 1156 (1996)

What follows deals (mostly) with statistical uncertainties!

How to make a more precise measurement?

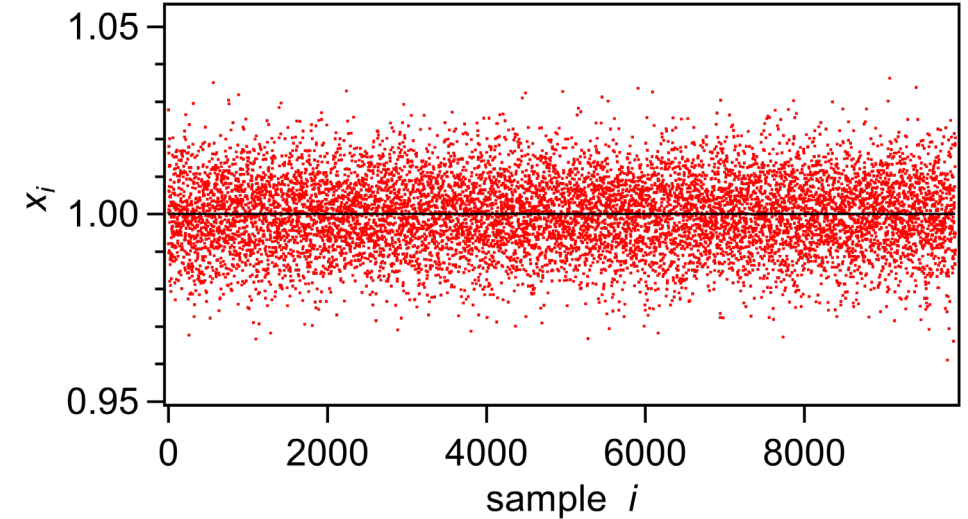
How to estimate the measurement uncertainty?

Characterization of a distribution

Single measurements x_i group around the “true” value μ

For an infinite number of measurements:

The distribution yields the parent distribution of x_i



Mean of the parent distribution

$$\mu = \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N x_i$$

Variance of the parent distribution (standard deviation σ)

$$\sigma^2 := \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N (x_i - \mu)^2$$

Characterization of a distribution

Transfer of events x_i into histogram:
Probability/frequency $P(x_j)$ of possible values x_j

Discrete distribution

$$\mu = \langle x \rangle = \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N x_i = \sum_j x_j P(x_j)$$

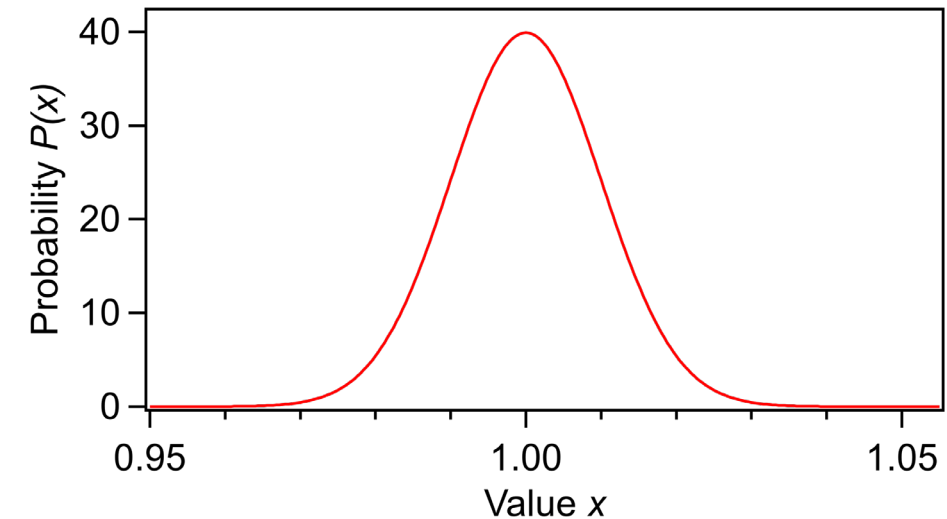
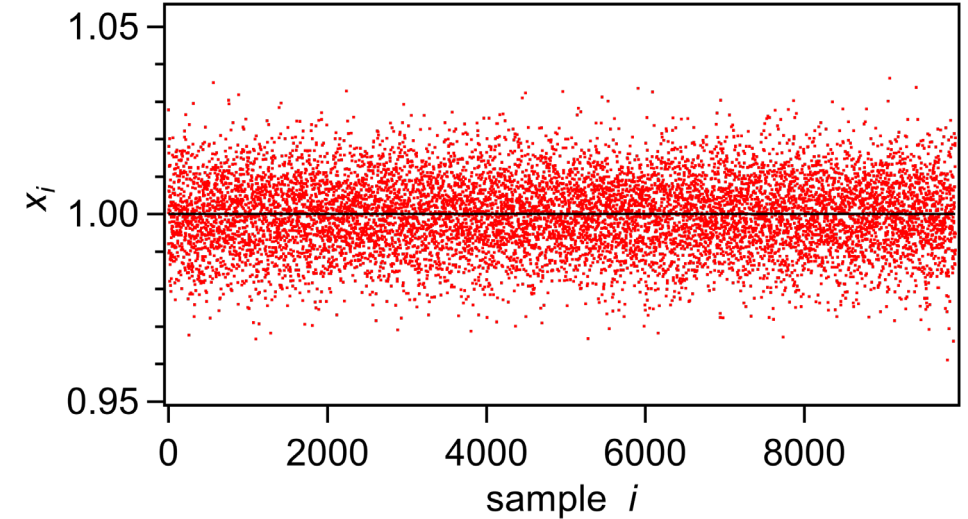
Continuous distribution

$$\mu = \int_{-\infty}^{+\infty} x' P(x') dx'$$

Normalization

$$\int_{-\infty}^{+\infty} P(x') dx' = 1$$

Problem: Parent distribution is in most cases not known!



Characterization of a sample distribution

For finite number of measurements:
The measurements are sample distributions
of the parent distribution

Mean of a sample distribution:

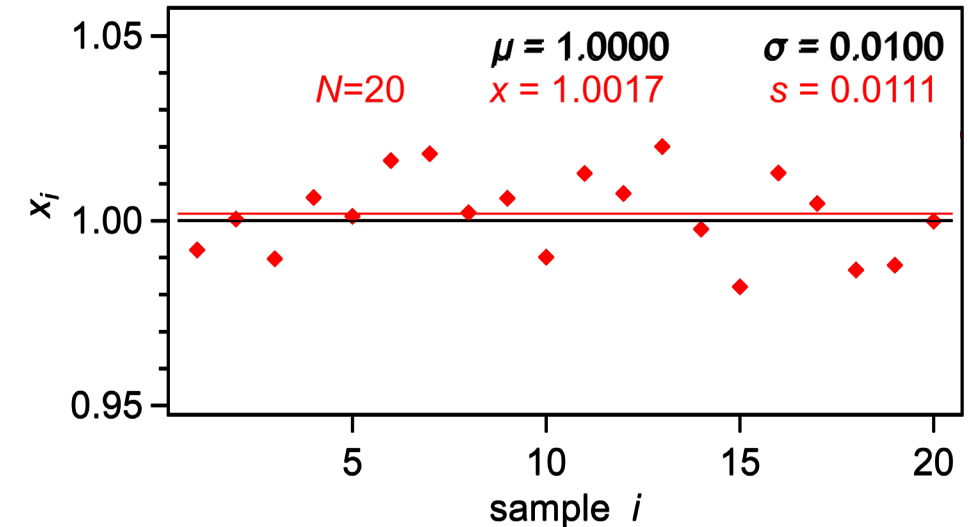
$$\bar{x} = \frac{1}{N} \sum_{i=1}^N x_i$$

Variance of a sample distribution

$$s^2 := \frac{1}{N-1} \sum_{i=1}^N (x_i - \bar{x})^2$$

Standard uncertainty of a sample distribution
(standard deviation of the mean)

$$u(x) := \sigma_{\bar{x}} := \frac{s}{\sqrt{N}} = \sqrt{\frac{1}{N(N-1)} \sum_{i=1}^N (x_i - \bar{x})^2}$$



Probability distributions

What is the probability of observing x in an experiment.

To answer this question we need to understand the parent distribution. Most common are:

Binomial distribution

$$P_B(x; n, p) = \binom{n}{x} p^x (1 - p)^{n-x}$$

Poisson distribution

$$P_B(x; \mu) = \frac{\mu^x}{x!} e^{-\mu}$$

Gaussian distribution

$$P_G(x; \mu, \sigma) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2}}$$

Lorentzian distribution

$$P_L(x; \mu, \Gamma) = \frac{1}{\pi} \frac{\Gamma/2}{(x-\mu)^2 + (\Gamma/2)^2}$$

Voigt = convolution of Lorentzian and Gaussian

Binomial distribution

Probability of observing in a random experiment x successes out of n tries when the probability of success is p

$$P_B(x; n, p) = \binom{n}{x} p^x (1 - p)^{n-x} = \frac{n!}{x!(n-x)!} p^x (1 - p)^{n-x}$$

Example: toss n coins, P_B = probability to find x coins with head up

p : probability head up

$1 - p$: probability head down (= tail up)

here: $p = 1/2$

Example: roll n dices, find x dices with a certain number, $p(x) = 1/6$, $(1 - p) = 5/6$



Binomial distribution (Dice rolling)

Example: Probability of rolling a “6” in 10 attempts

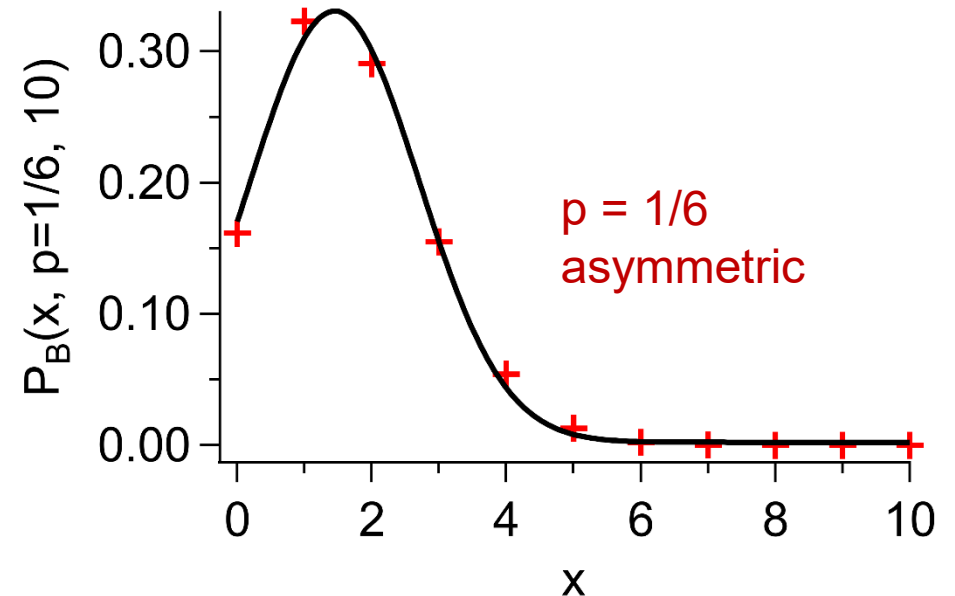
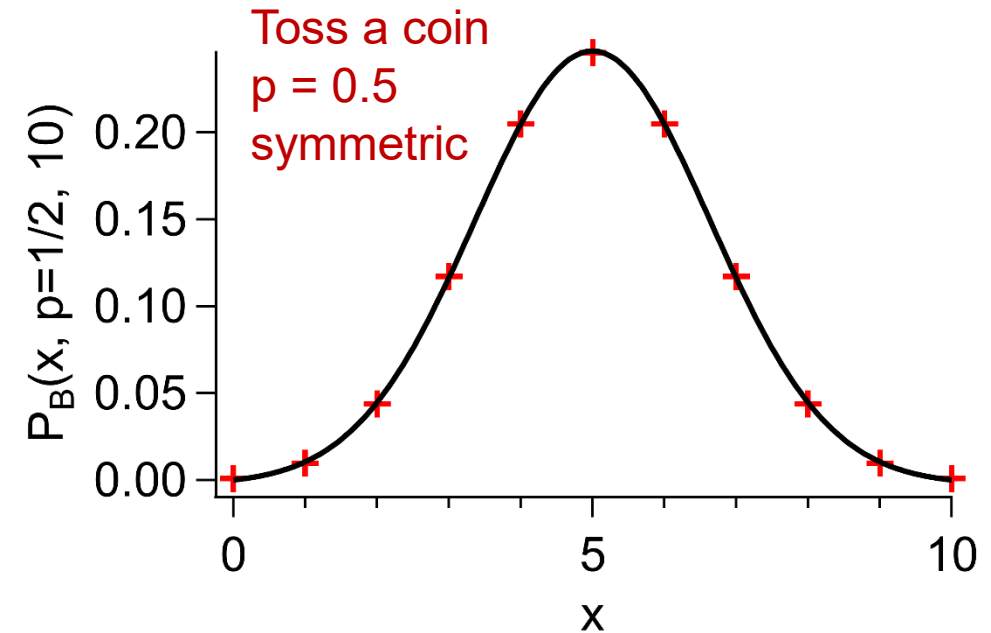
Probability of observing in a random experiment x successes out of n tries when the probability of success is p

$$\begin{aligned}P_B(x; n, p) &= \binom{n}{x} p^x (1 - p)^{n-x} \\&= \frac{n!}{x!(n-x)!} p^x (1 - p)^{n-x}\end{aligned}$$

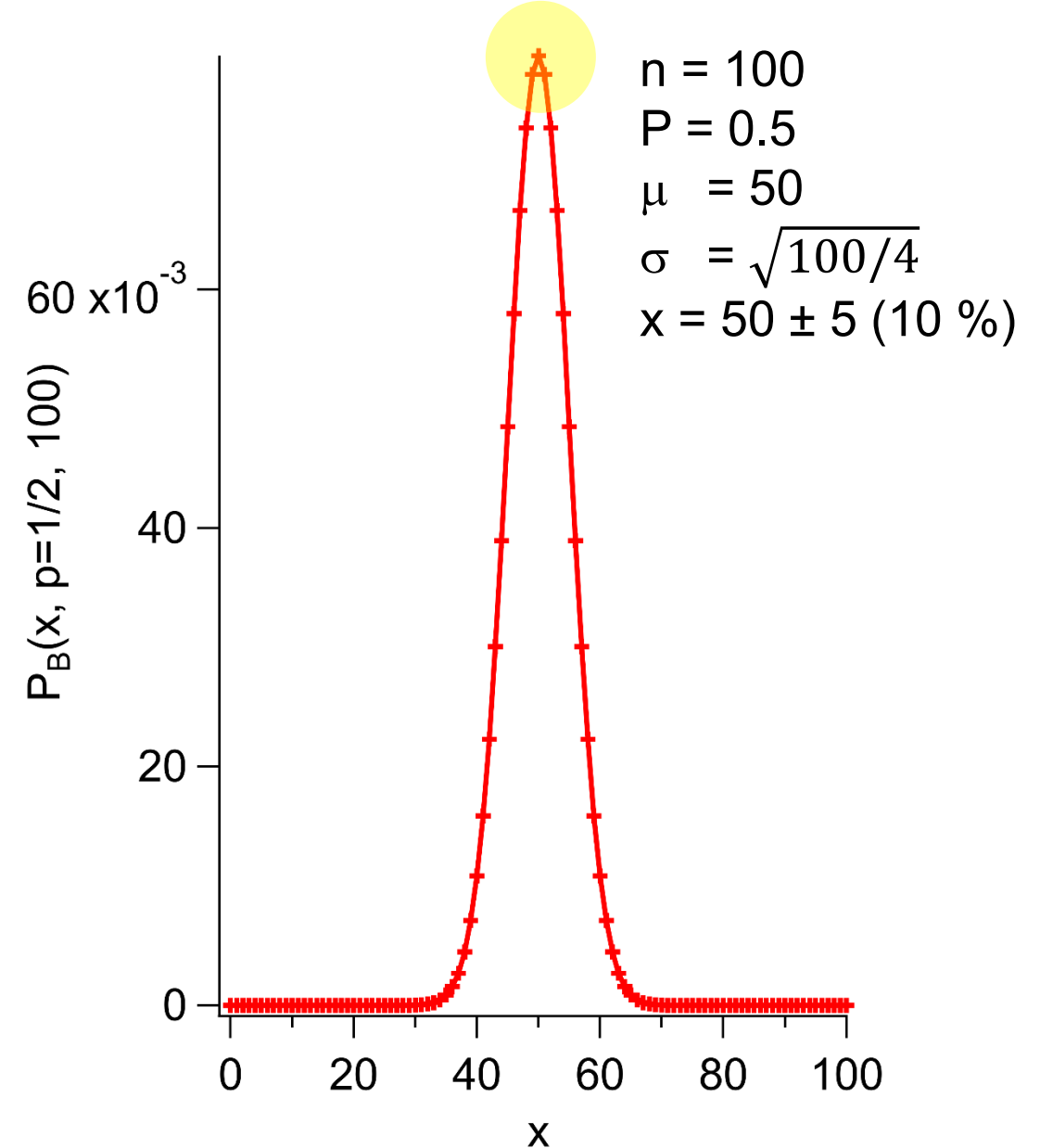
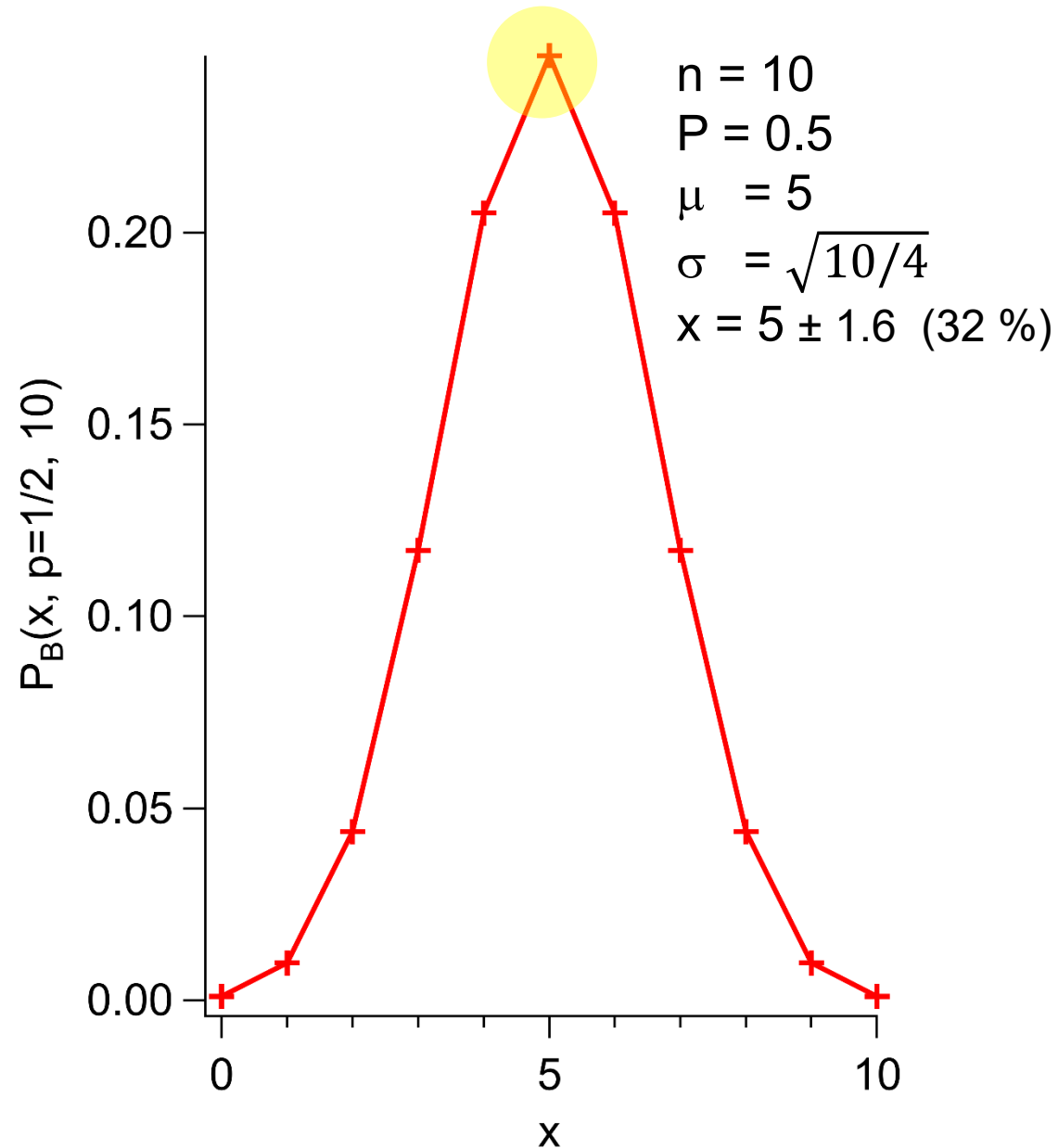
Mean: $\mu = np$

Variance: $\sigma^2 = np(1 - p)$

10 dices, $p = 1/6$, $\mu(6) = 10/6 = 1.67$, $\sigma^2 = 50/36$, $\sigma = 1.17$



For large n and small steps Δx , evaluation becomes impractical !



Poisson distribution: Counting statistics

For small probabilities of success, $p \ll 1$, the Binomial distribution can be approximated by

the **Poisson distribution**:

$$P_P(x; \mu) = \frac{\mu^x}{x!} e^{-\mu}$$

Mean

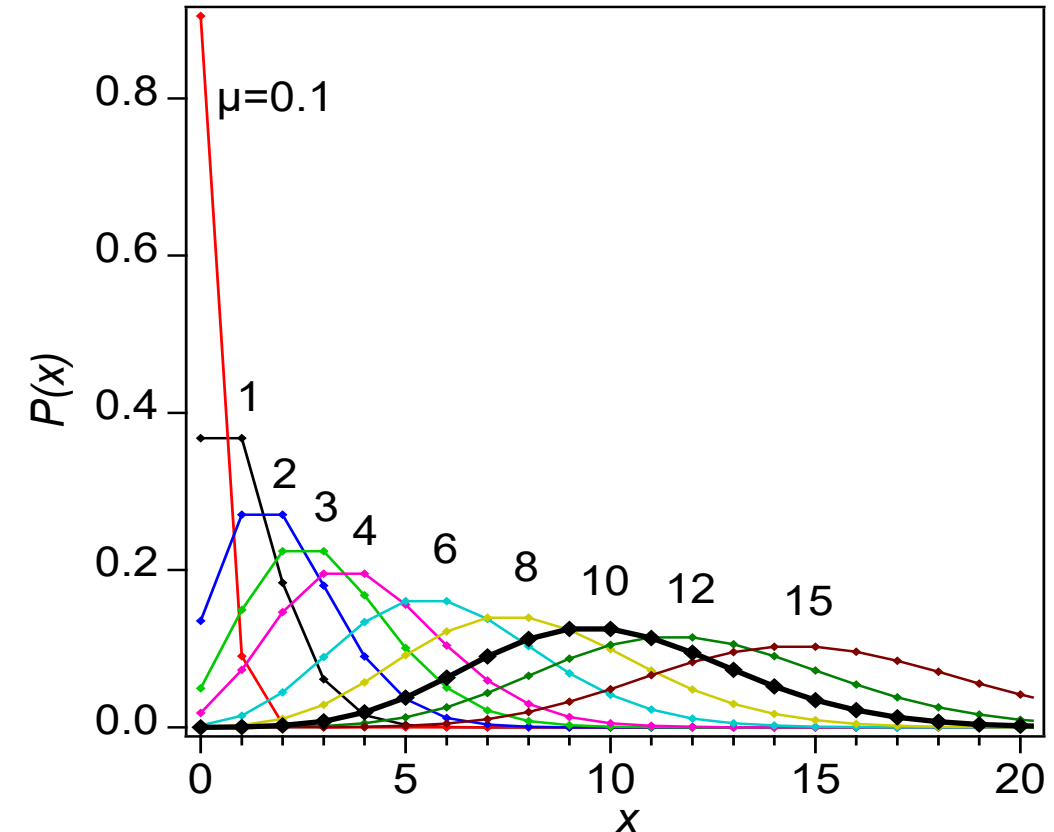
$$\bar{x} = \sum_{x=0}^{\infty} x \frac{\mu^x}{x!} e^{-\mu} = \mu$$

Variance

$$\sigma^2 = \sum_{x=0}^{\infty} (x - \mu)^2 \frac{\mu^x}{x!} e^{-\mu} = \mu$$

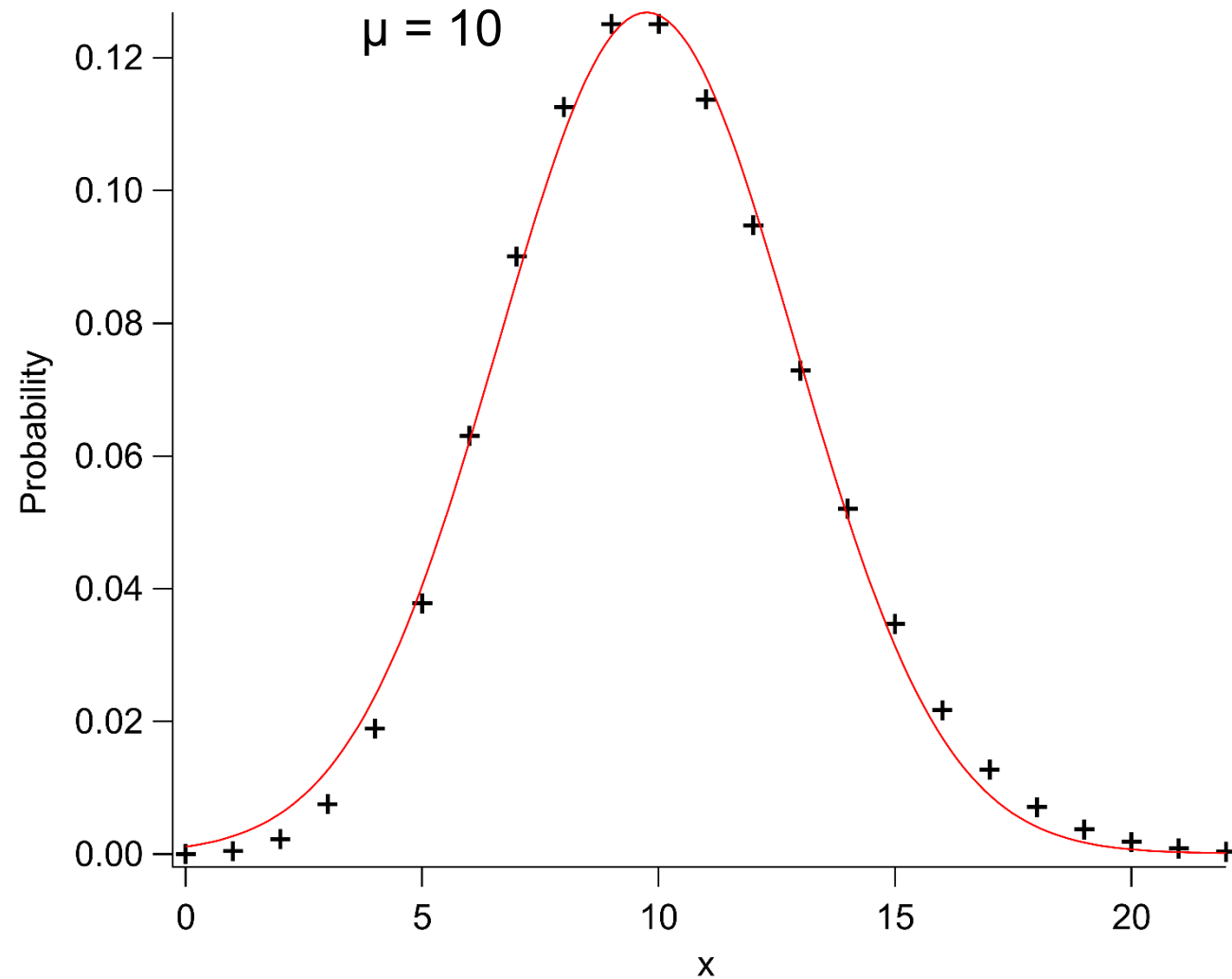
$$\frac{\sigma}{\mu} = \frac{1}{\sqrt{\mu}}$$

Poisson distribution
only depends on μ



- Recap of previous seminar
- Probability distribution: Gaussian and Lorentzian
- Error propagation
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- How to report and present results

Poisson vs. Gaussian distribution



For $\mu > 10$, the Poisson distribution (+) is well approximated by a **Gaussian** distribution.

Gaussian distribution

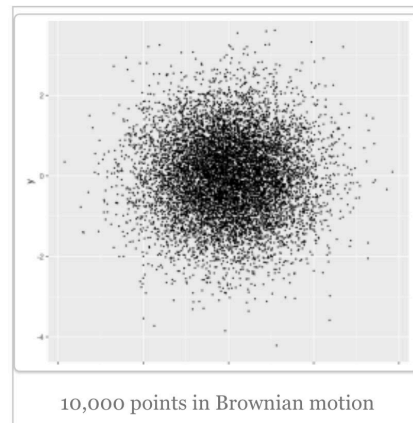
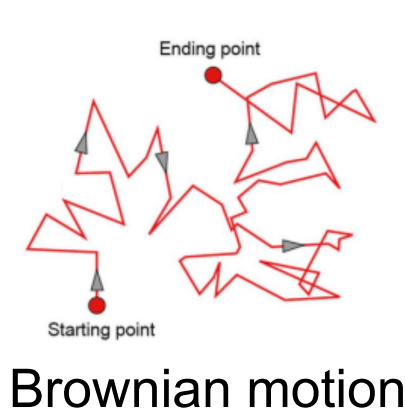
For $\mu \gg 1$ the Poisson distribution is nicely approximated by a **Gaussian distribution**.

“There are several derivations of the Gaussian distribution from first principles, none of them as convincing as the fact that the distribution is **reasonable**, that it has a fairly **simple analytic form**, and that it is accepted by convention and experimentation to be the **most likely distribution for most experiments**.

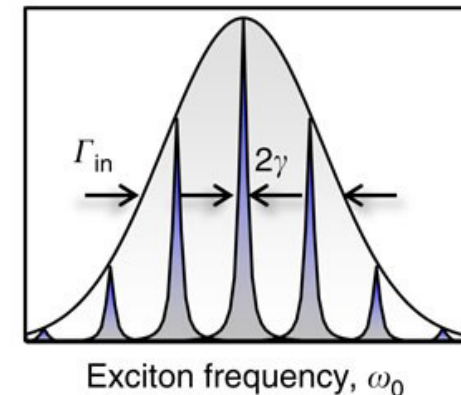
In addition, it has the satisfying characteristic that the **most probable estimate of the mean μ** from a random sample of observations x is the **average** of those observations x .”

(From: P. R. Bevington, and D. K. Robinson, “Data Reduction and Error Analysis for the Physical Sciences”)

Examples:



Inhomogeneous broadening

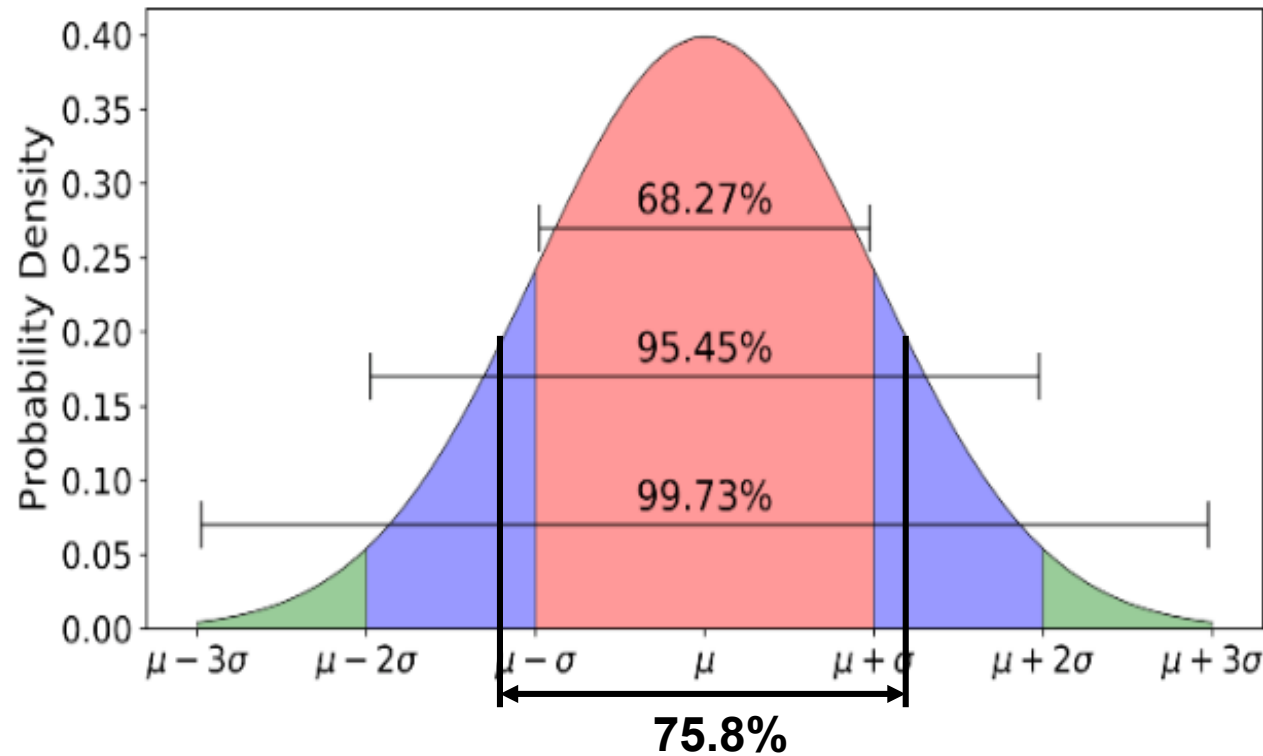


Moody et al.
Nat. Commun.
6, 8315 (2015)

Gaussian distribution

Probability of one measurement being within $[-\sigma, +\sigma]$

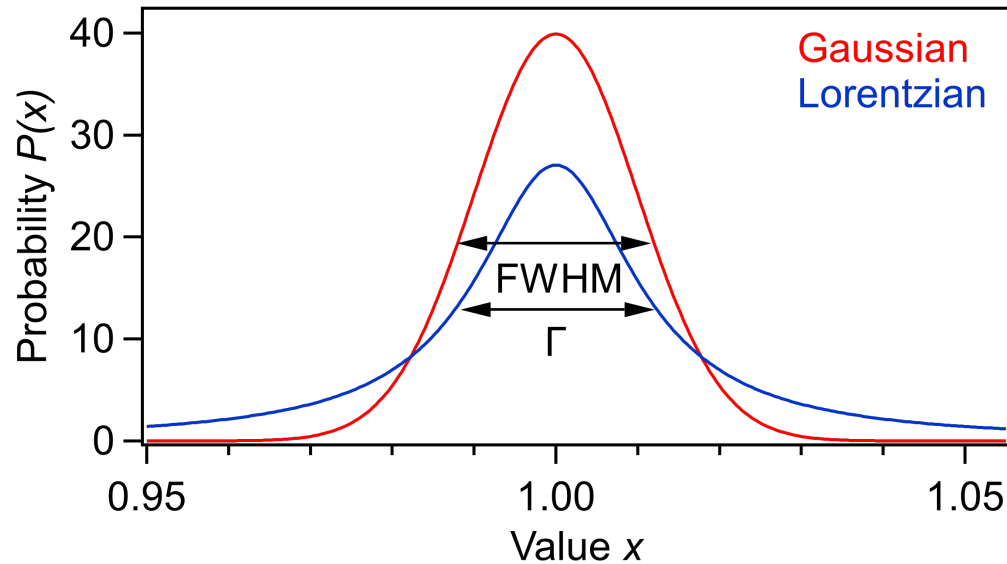
$$P_G(x \text{ within } \sigma) = \int_{\mu-\sigma}^{\mu+\sigma} dx \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2}}$$



$$\text{FWHM} = 2\sigma\sqrt{2\ln(2)} \approx 2.354\sigma$$

Lorentzian distribution (Cauchy distribution)

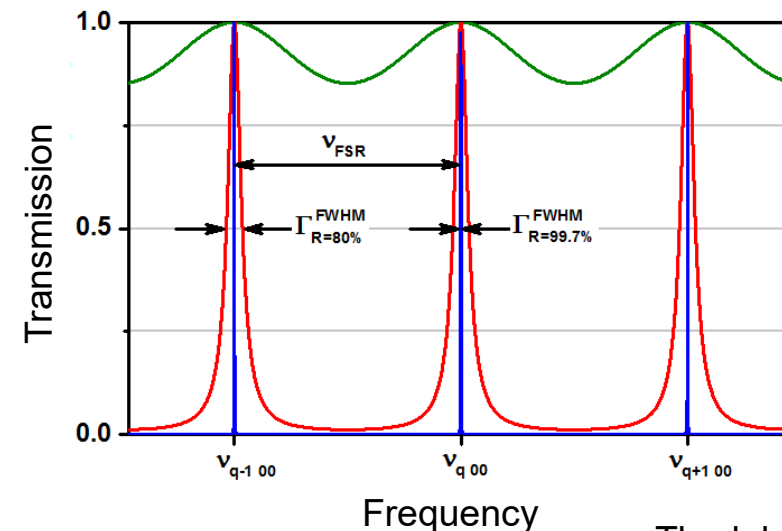
$$P_L(x; \mu, \Gamma) = \frac{1}{\pi} \frac{\Gamma/2}{(x-\mu)^2 + (\Gamma/2)^2}$$



Standard deviation not defined,
because of strong “wings”.

Lorentzian distribution:
Fourier transform of exponential decay

- ➔ Width of frequency distribution is related to damping:
- Finite lifetime of excited states
 - Lineshape of laser cavity modes due to losses (HeNe laser, Fabry-Perot)

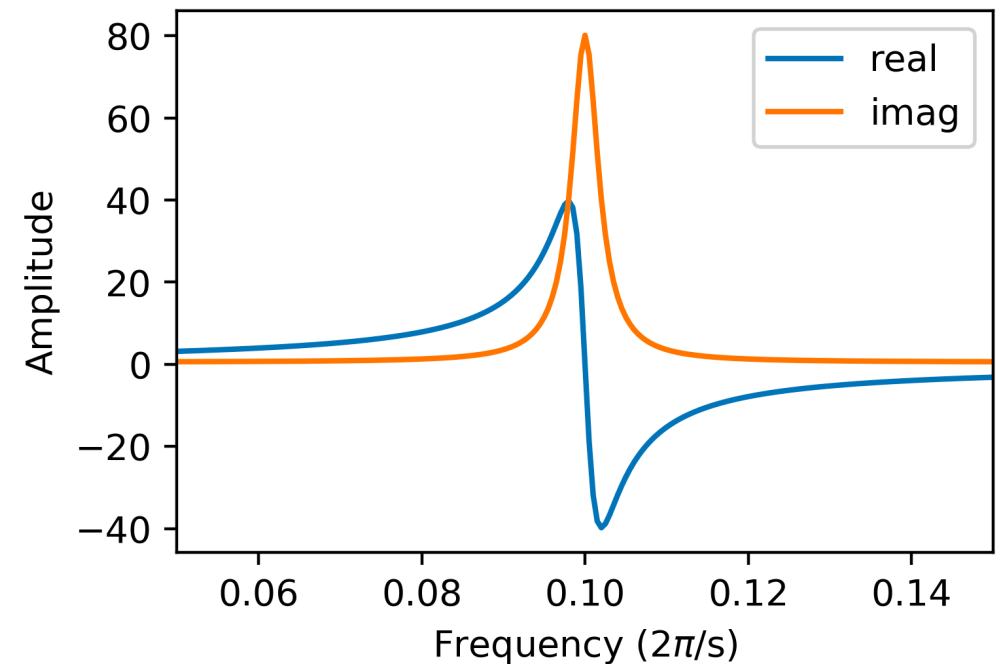
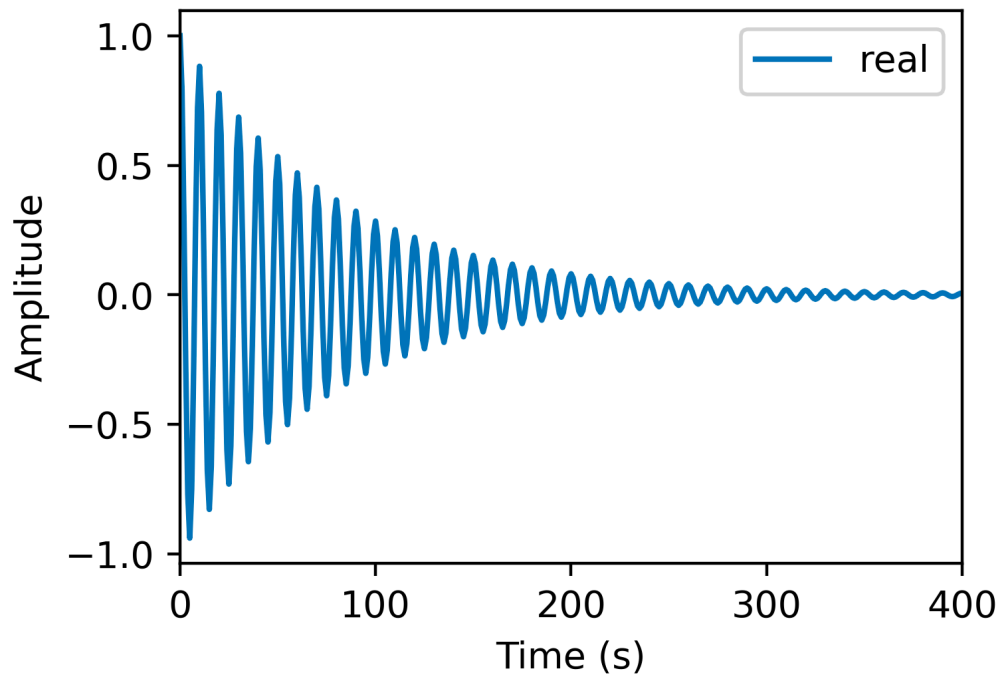


Thorlabs: Fabry-Perot tutorial

Damped oscillator (light-wave emission, dipole radiation)

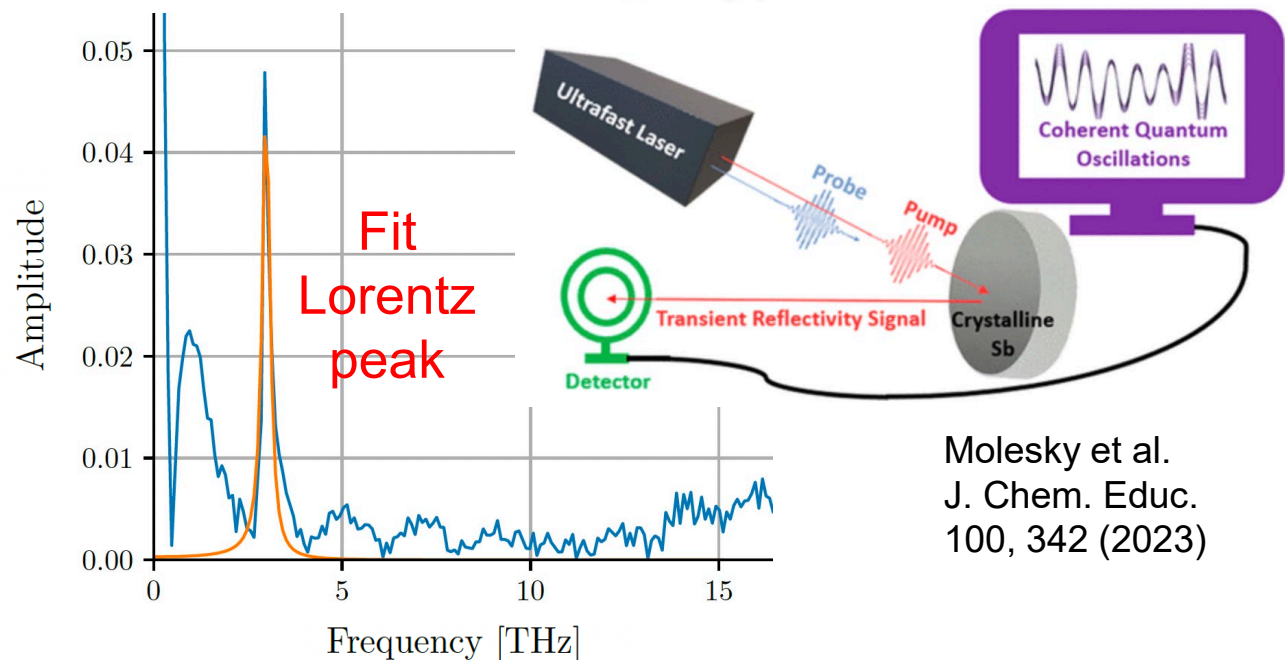
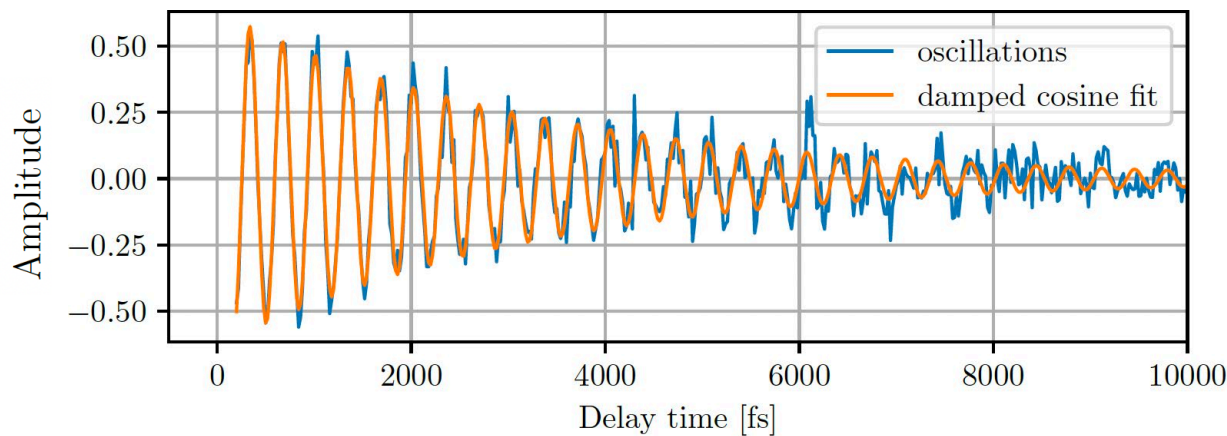
$$f(t) = e^{i\omega_0 t} e^{-\gamma t}$$

$$\mathcal{F}[f(t), \omega] \propto \frac{\omega_0 - \omega}{(\omega_0 - \omega)^2 + \gamma^2} + i \frac{\gamma}{(\omega_0 - \omega)^2 + \gamma^2}$$

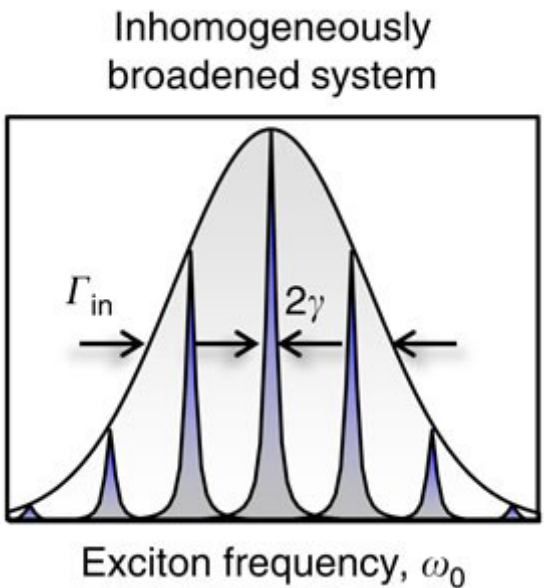
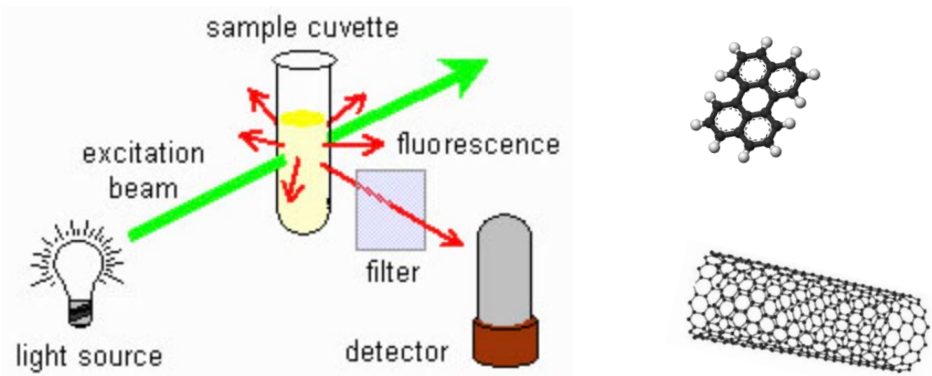


Examples of Lorentz vs Gauss Line Shape

Ma16: Coherent phonons



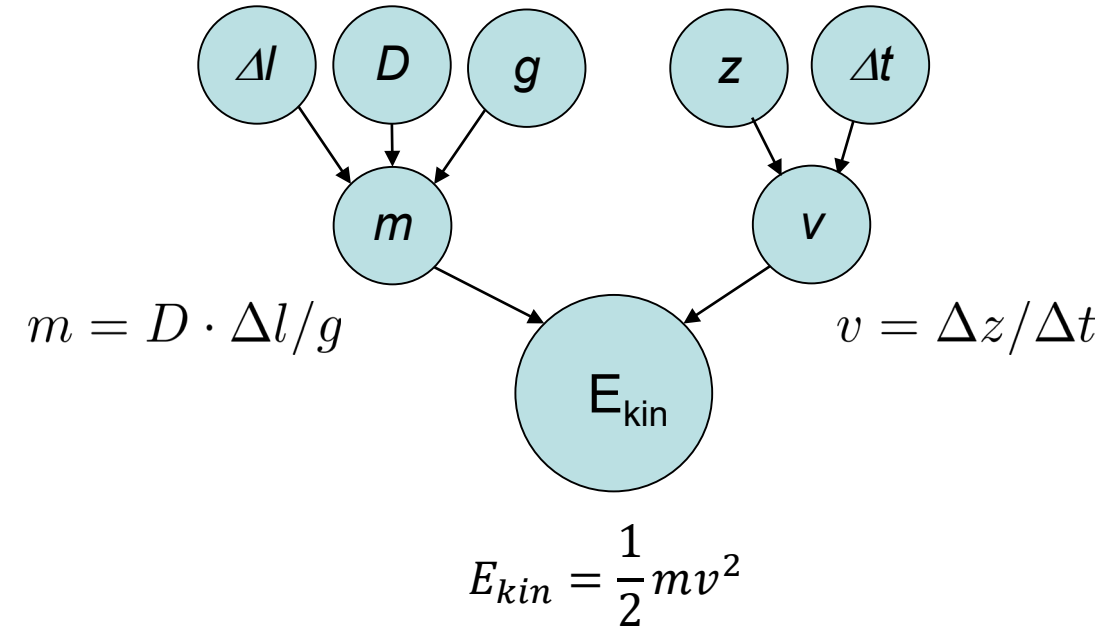
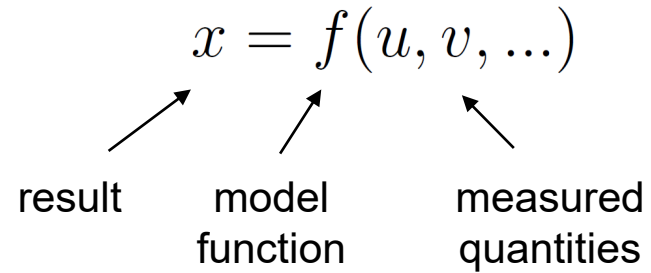
Ma9: Photoluminescence



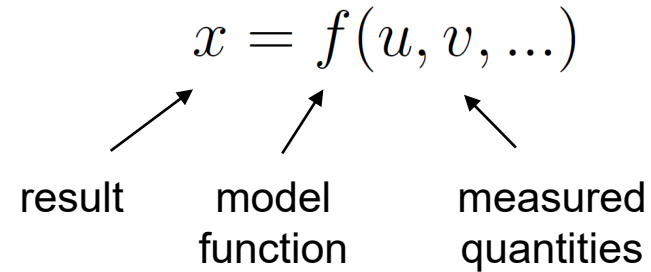
Moody et al. Nat. Commun. 6, 8315 (2015)

- Recap of previous seminar
- Probability distribution: Gaussian and Lorentzian
- Error propagation
- Data fitting
- How to report and present results

Functional dependences of quantities



- Test model by systematically varying parameters
- Measuring $f(u, v, \dots; a, b, \dots)$
Determine model parameters a, b, \dots by a fit
- What is the uncertainty of the result?



Data evaluation procedures:

a.) Calculate individual $x_i = f(u_i, v_i, \dots)$ ➤ form histogram $h(x_j)|dx_j|$

➤ If u_i and v_i belong together

b.) Calculate means and take most probable values $\bar{x} = f(\bar{u}, \bar{v}, \dots)$

➤ If u_i and v_i from independent measurements

➤ Error propagation?

Propagation of uncertainty

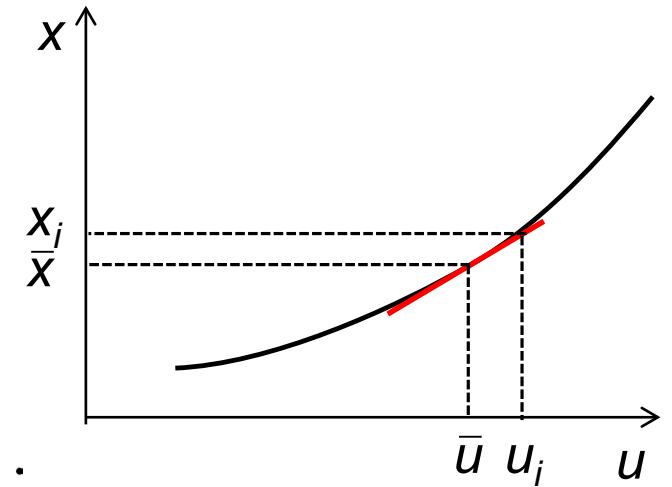
For measured quantities u, v, \dots

$$x = f(u, v, \dots)$$

Assuming that the deviations in u_i, v_i, \dots are small, deviations in x_i can be expressed as

Taylor expansion $(x_i - \bar{x}) = \left. \frac{\partial f}{\partial u} \right|_{\bar{u}} (u_i - \bar{u}) + \left. \frac{\partial f}{\partial v} \right|_{\bar{v}} (v_i - \bar{v}) + \dots$

(derivatives evaluated for variables fixed at mean)

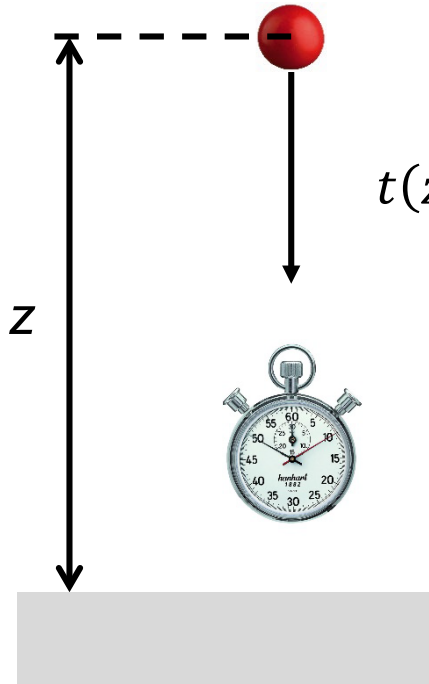


Variance:

$$s_f^2 = \frac{1}{N-1} \sum_i (x_i - \bar{x})^2 = \frac{1}{N-1} \sum_i \left[(u_i - \bar{u}) \left(\frac{\partial f}{\partial u} \right) + (v_i - \bar{v}) \left(\frac{\partial f}{\partial v} \right) + \dots \right]^2$$
$$s_f^2 = s_u^2 \left(\frac{\partial f}{\partial u} \right)^2 + s_v^2 \left(\frac{\partial f}{\partial v} \right)^2 + 2s_{uv} \left(\frac{\partial f}{\partial u} \right) \left(\frac{\partial f}{\partial v} \right) + \dots$$

Propagation of uncertainty

$$s_f^2 = s_u^2 \left(\frac{\partial f}{\partial u} \right)^2 + s_v^2 \left(\frac{\partial f}{\partial v} \right)^2 + 2s_{uv}^2 \left(\frac{\partial f}{\partial u} \right) \left(\frac{\partial f}{\partial v} \right) + \dots$$



$$t(z_0, t_0) = \sqrt{\frac{2(z + z_0)}{g}} + t_0$$

$$\frac{\partial t}{\partial z_0} < 0, \quad \frac{\partial t}{\partial t_0} > 0$$

Covariance: characterizing the correlation between deviations in different variables

$$s_{uv}^2 = \frac{1}{N-1} \sum (u_i - \bar{u})(v_i - \bar{v})$$

- can be positive or negative
- vanishes for uncorrelated quantities

Propagation of standard uncertainty

$$\sigma_{\bar{x}}^2 = \frac{s_f^2}{N} = \sigma_{\bar{u}}^2 \left(\frac{\partial f}{\partial u} \right)^2 + \sigma_{\bar{v}}^2 \left(\frac{\partial f}{\partial v} \right)^2 + \cancel{2\sigma_{\bar{u}\bar{v}}^2 \left(\frac{\partial f}{\partial u} \right) \left(\frac{\partial f}{\partial v} \right)} + \dots$$

for uncorrelated errors
because

$$s_{uv}^2 = \frac{1}{N-1} \sum (u_i - \bar{u})(v_i - \bar{v}) = 0$$

Often it is sufficient to use just the biggest term!

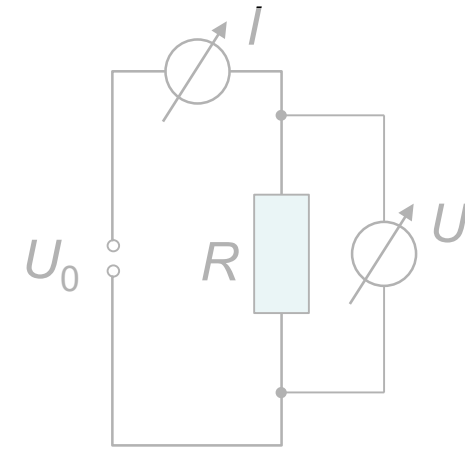
Propagation of uncertainty: *examples*

$$\sigma_{\bar{x}}^2 = \frac{s_f^2}{N} = \sigma_{\bar{u}}^2 \left(\frac{\partial f}{\partial u} \right)^2 + \sigma_{\bar{v}}^2 \left(\frac{\partial f}{\partial v} \right)^2$$

Resistance measurement:

Measure current I and voltage U on resistor to determine resistance R

$$R = \frac{U}{I} \quad \sigma_R \approx \sqrt{\sigma_U^2 \left(\frac{1}{I} \right)^2 + \sigma_I^2 \left(\frac{U}{I^2} \right)^2}$$



Conversion of wavelength to energy:

$$E \text{ [eV]} = \frac{hc}{\lambda} \approx \frac{1240}{\lambda \text{ [nm]}} \text{ eV} \quad \lambda = (635 \pm 6) \text{ nm}$$
$$\rightarrow E = (? \pm ?) \text{ eV}$$

- Recap of previous seminar
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Comparing the assumed probability distribution function P to a measured histogram $h(x_j)$:

$$\chi^2 = \sum_{j=1}^n \frac{[h(x_j) - NP(x_j)]^2}{\sigma_j^2(h)}$$

Expected for pure counting statistics:

$$\langle \chi^2 \rangle = \nu = \underset{\substack{\uparrow \\ \text{\# data/bins}}}{n} - \underset{\substack{\uparrow \\ \text{\# constraints}}}{n_c}$$

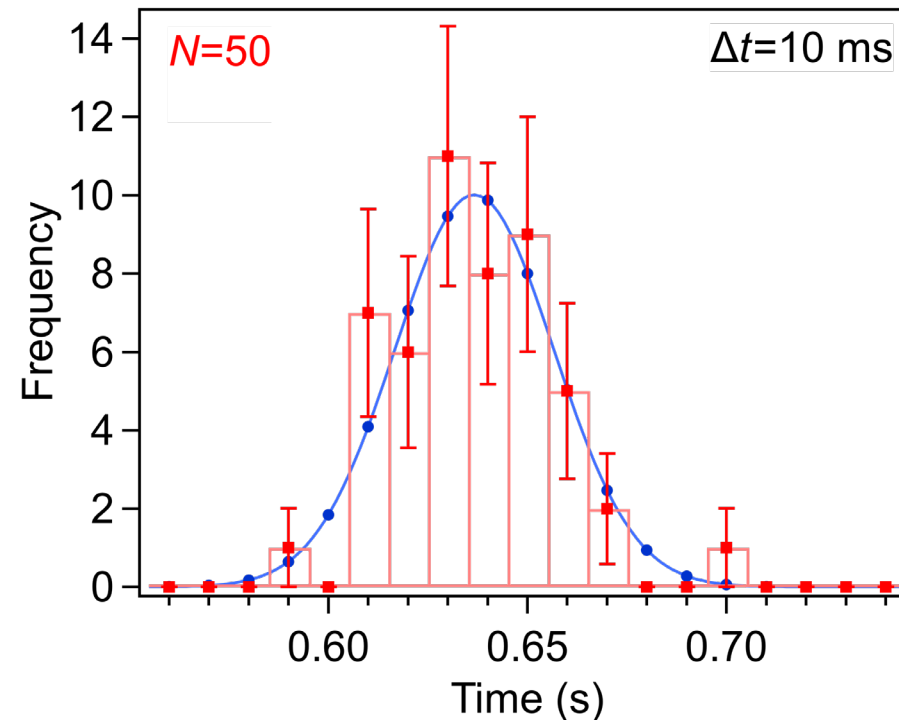
Reduced chi-squared:

$$\langle \chi_\nu^2 \rangle = \left\langle \frac{\chi^2}{\nu} \right\rangle = 1$$

$n = 9$ bins

$n_c = 3$

Gaussian: A, \bar{x}, σ



$$\chi^2 = 3.27$$

$$\nu = 9 - 3 = 6$$

$$\chi_\nu^2 = 0.55$$

Least-squares fit to a line

Fit function:

$$y(x) = a + bx$$

Data set of (x_i, y_i) with uncertainties $\sigma_{y_i} = \sigma_i$

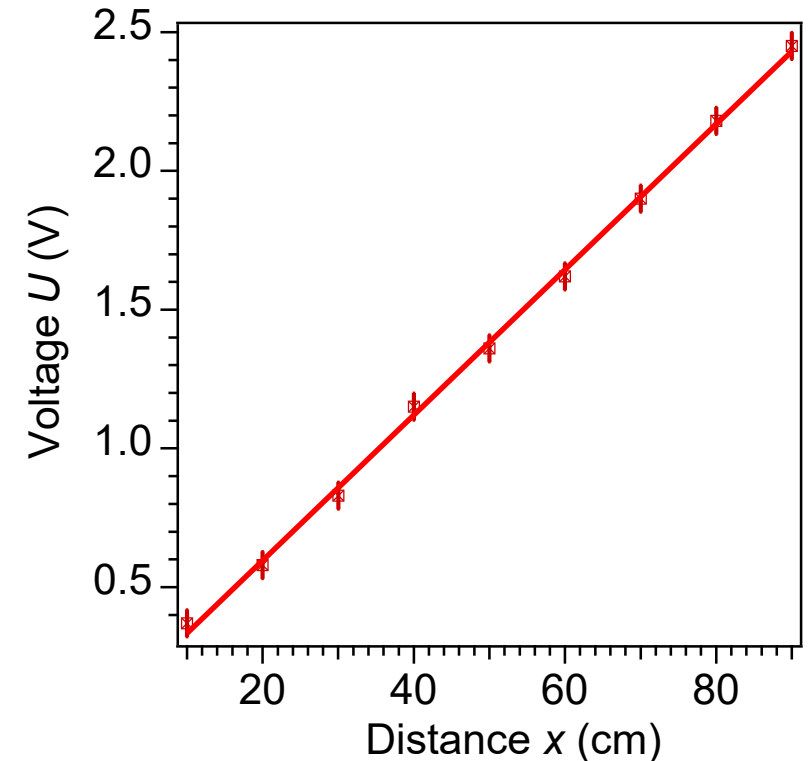
Assuming Gaussian uncertainties, the most probable parameters a, b are obtained at minimum of

$$\chi^2 = \sum_{i=1}^N \frac{[y_i - y(x_i)]^2}{\sigma_i^2} = \sum_{i=1}^N \left[\frac{1}{\sigma_i} (y_i - a - bx_i) \right]^2 \approx \nu$$

Reduced chi-square

$$\chi_\nu^2 = \frac{\chi^2}{\nu} \approx 1 \quad \text{with } \nu = N - N_c \quad (\text{here } N_c = 2)$$

good fit



Chi-square

$$\chi^2 = \sum_{i=1}^N \frac{[y_i - y(x_i)]^2}{\sigma_i^2}$$

Reduced chi-square

$$\chi^2_\nu = \frac{\chi^2}{\nu} \approx 1$$

- *Takes uncertainties into account*
- *Test whether data follow particular distribution*

R squared

(coefficient of determination)

$$R^2 = 1 - \frac{\sum_{i=1}^N [y_i - y(x_i)]^2}{\sum_{i=1}^N (y_i - \bar{y})^2} \approx 1$$

Fit function

Mean of data

for good fit

- *Only from residuals*
- *Measure for the variation of the data compared to model*

Least-squares fit to a line: Linear regression

Analytical solution:

$$a = \frac{1}{\Delta} \left(\sum \frac{x_i^2}{\sigma_i^2} \sum \frac{y_i}{\sigma_i^2} - \sum \frac{x_i}{\sigma_i^2} \sum \frac{x_i y_i}{\sigma_i^2} \right)$$

$$b = \frac{1}{\Delta} \left(\sum \frac{1}{\sigma_i^2} \sum \frac{x_i y_i}{\sigma_i^2} - \sum \frac{x_i}{\sigma_i^2} \sum \frac{y_i}{\sigma_i^2} \right)$$

$$\text{with } \Delta = \sum \frac{1}{\sigma_i^2} \sum \frac{x_i^2}{\sigma_i^2} - \left(\sum \frac{x_i}{\sigma_i} \right)^2$$

Variance of fit parameters:

$$\sigma_a^2 = \frac{1}{\Delta} \sum \frac{x_i^2}{\sigma_i^2} \quad \sigma_b^2 = \frac{1}{\Delta} \sum \frac{1}{\sigma_i^2}$$

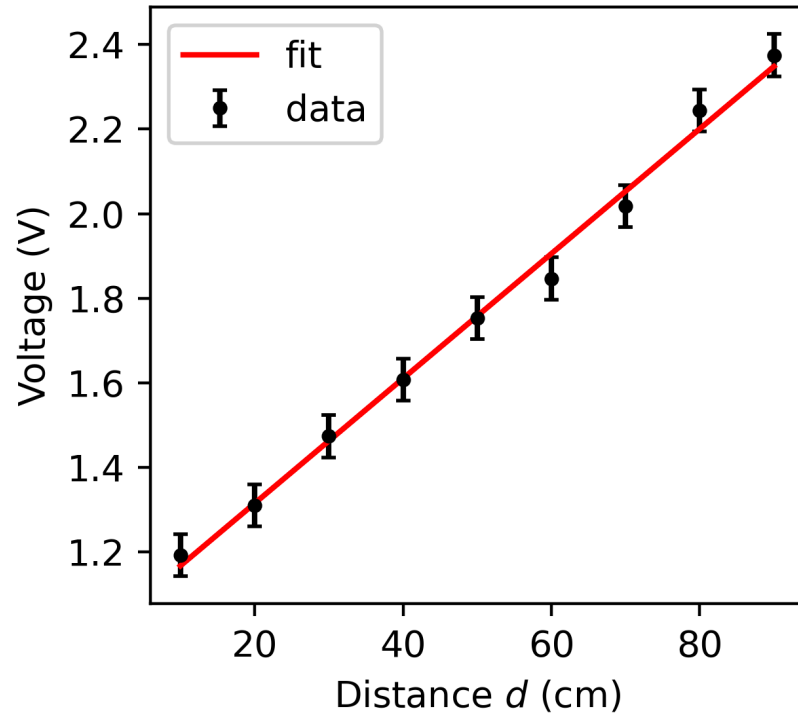
Estimation of uniform variance

$$\sigma'^2 \simeq s^2 = \frac{1}{N-2} \sum (y_i - a - bx_i)^2$$

Example 1: Linear fit

Potential drop along a wire

$$U(d) = RI d/d_0$$



$$\sigma_U = 0.05 \text{ V}$$

$R^2 = 0.994 \rightarrow$ good fit

$$\chi^2 = 3.21 \quad \text{for } N=9, \nu=7$$

- χ^2 is too small.
- Standard deviation was overestimated

Corrected estimate:

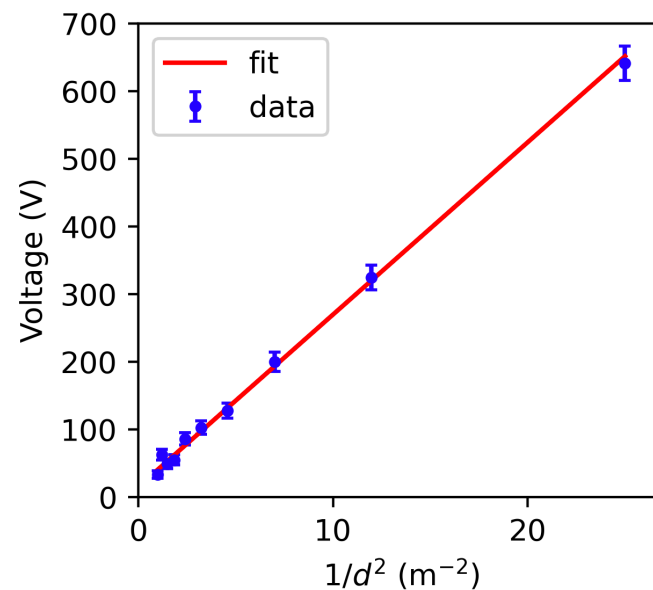
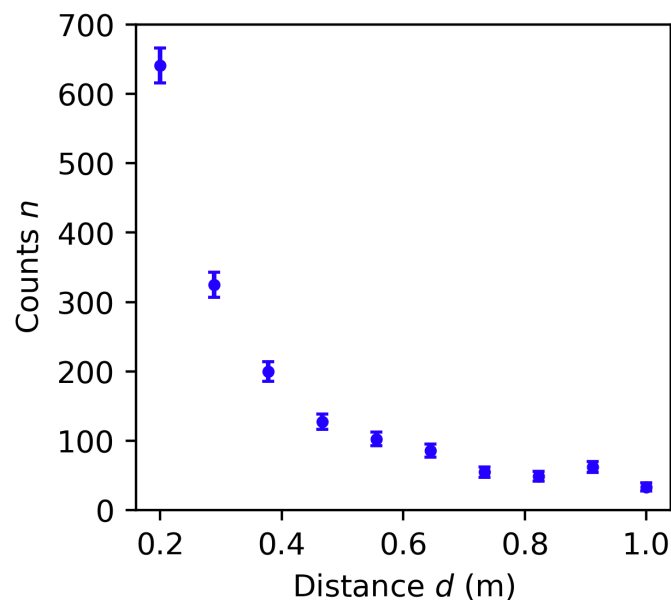
$$\sigma' \approx \sigma \sqrt{\chi_\nu} = 0.033 \text{ V}$$

Note: Standard deviation concerns only the **random errors**.

Total uncertainty due to **systematic errors** might be larger!

Example 2: Linearize data

Radioactive source at distance d



$$R^2 = 0.998$$

$$\chi^2 = 9.3 \quad \text{for } N=10, \nu=8$$

➤ Estimation of σ_n fits well.

$I(d) = I_0/d^2 + I_{\text{dark}}$ ➡ Fit allows separation of signal and background

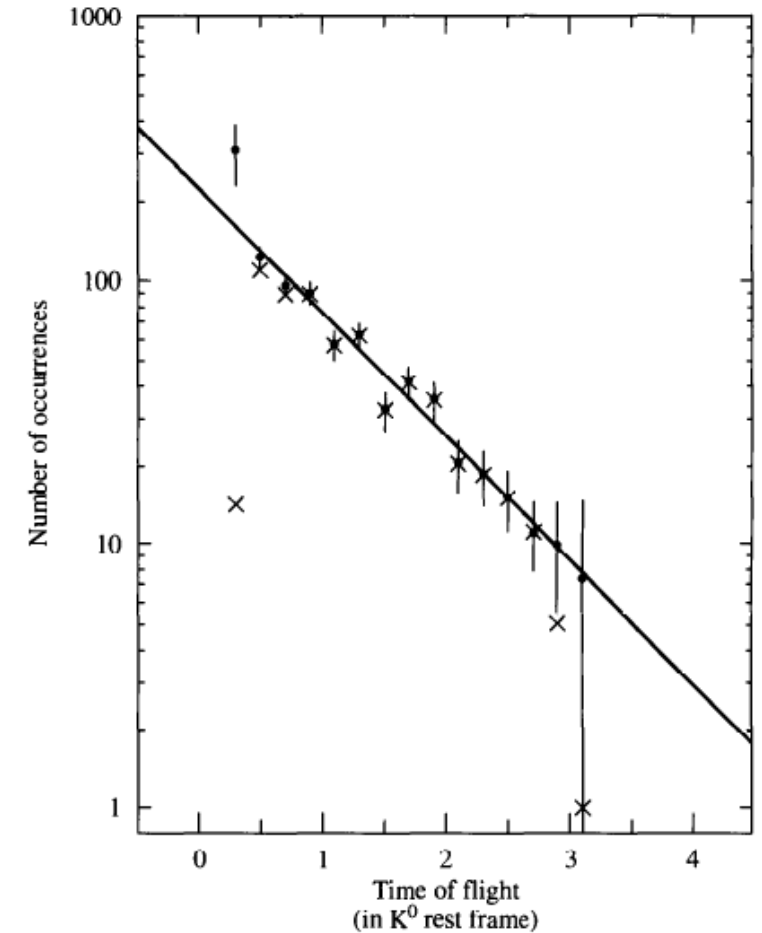
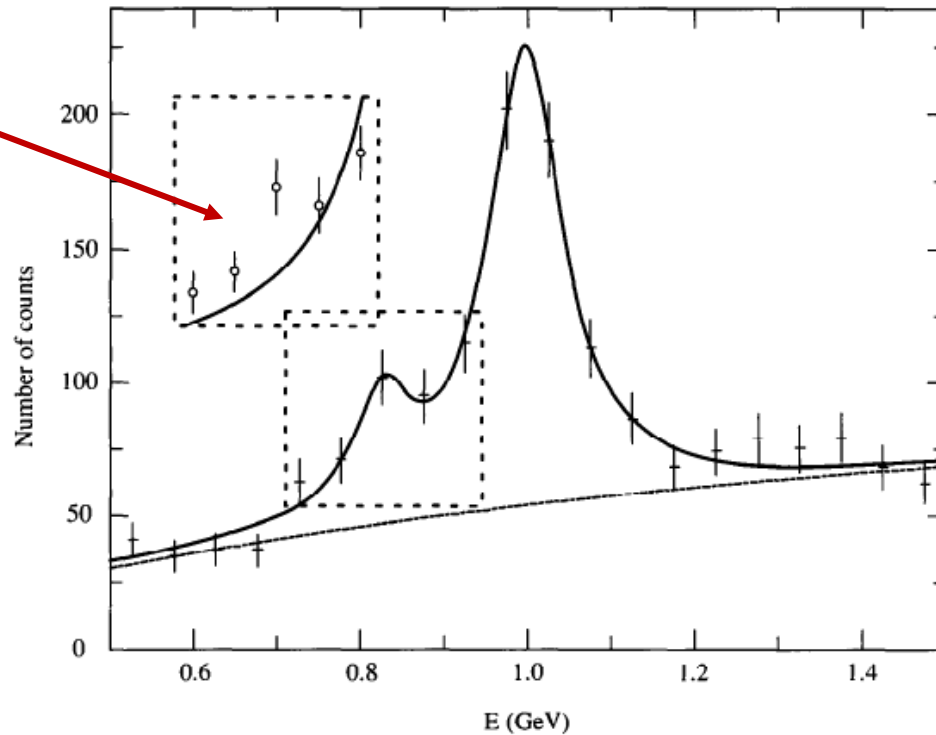
$$\sigma_n = \sqrt{n}$$

Non-linear fitting

If fitting function is non-linear in parameters a , b , ...

- try to transform to linear fit
- numerical least-mean-squares fit

Fit with
one peak



“A fit can never be better than your model function.”

- Recap of previous seminar
- Probability distribution: Gaussian and Lorentzian
- Error propagation
- Data fitting
- How to report and present results

Fluctuations in reading of physical instrument (limited instrumental precision)

Often independent of the actual value
e.g. length, mass, voltage, current...



- Refer to instrumentation manuals or other test measurements
(*external estimate*)
- Estimate of standard deviation from repeated measurements
(*internal estimate*)

Counting experiments:

x represents the count rate in a detector

Statistical fluctuations for finite sample

- Uncertainty can be deduced from data
- Poisson statistics! $\sigma = \sqrt{\mu}$

How to report measurement results and uncertainties

- for $N = 1$ measurement, report instrumental uncertainty

- for $N > 1$ measurement, report

$$x = \bar{x} \pm \sigma_{\bar{x}}$$

$$\text{with } \sigma_{\bar{x}} = \sqrt{\frac{1}{N(N-1)} \sum_{i=1}^N (x_i - \bar{x})^2}$$

(or from fit)

- For fit also report reduced chi-squared and R^2 value.
- Indicate estimated *systematic* uncertainty in addition.
- Report uncertainty with 2 significant digits if intermediate result, and 1 digit for final result.
- Report result with same precision as uncertainty.

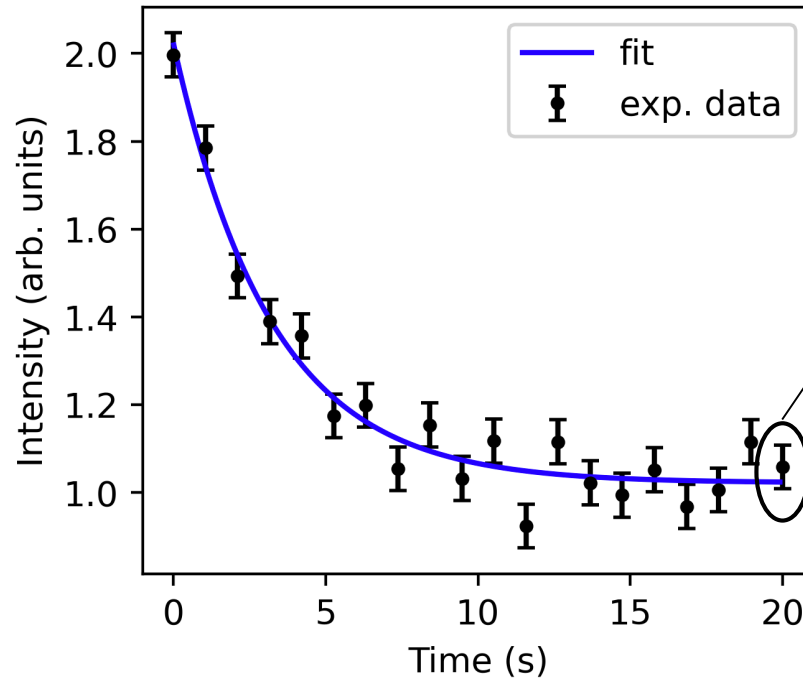
How to state measurement results and uncertainties

Exercise: State the quantity with uncertainty – convert to scientific notation if helpful

quantity	standard uncertainty	
50.003 m	0.1m	$(50.0 \pm 0.1) \text{ m}$
53612 s	300 s	$(5.36 \pm 0.03) \cdot 10^4 \text{ s}$
632.497 nm	0.153 nm	$(632.5 \pm 0.2) \text{ nm}$
367134 C	783 C	$(3.671 \pm 0.008) \cdot 10^5 \text{ C}$

How to plot uncertainties of measured data

For small data set:

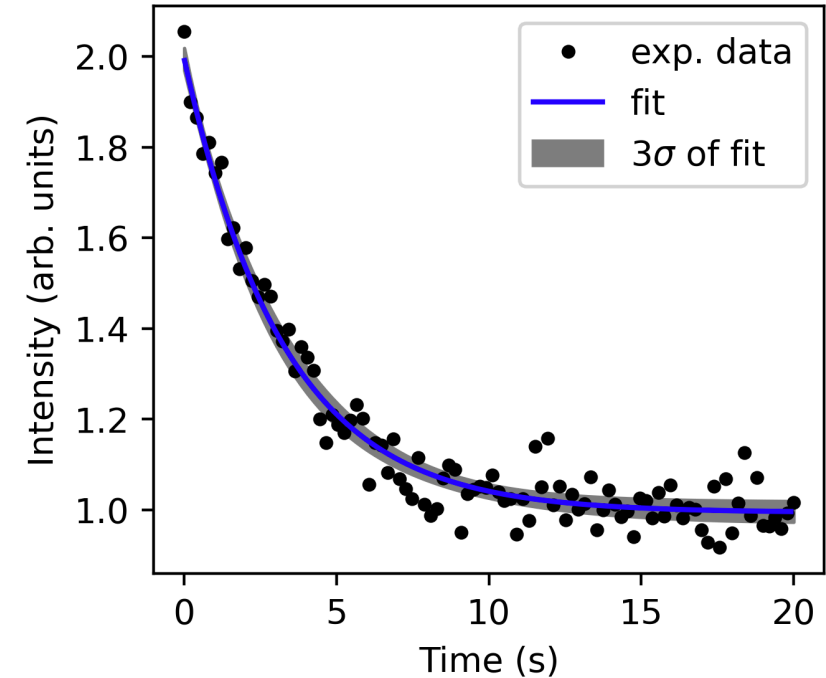


Error bar:

$2\sigma_{\bar{x}}$

- Show error bars
- Error bars show two standard deviations

For large data set:



- Plot uncertainty of fit
- Or use binning

Take home...

- Keep good experimental records
- Try to measure in several ways
- Calibrate instruments, read their manuals
- Check rounding errors
- Be mindful of largest uncertainty source
- The uncertainty of a fit result does not consider systematic deviations
- **Never report measurement results without uncertainties**

- P. R. Bevington, and D. K. Robinson, *“Data Reduction and Error Analysis for the Physical Sciences”*
- P. Möhrke, and B.-U. Runge, *“Arbeiten mit Messdaten”*
- *“Guide to uncertainty propagation and Error analysis”* Stony Brook U.
- *“A beginner guide to uncertainty of measurement”* by Stephanie Bell
- Slides and python code examples will be uploaded to website

Examples with Python